

RESEARCH ARTICLE

The Effect of Inadequate Initial Empiric Antimicrobial Treatment on Mortality in Critically Ill Patients with Bloodstream Infections: A Multi-Centre Retrospective Cohort Study

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Abstract

Hospital mortality rates are elevated in critically ill patients with bloodstream infections. Given that mortality may be even higher if appropriate treatment is delayed, we sought to determine the effect of inadequate initial empiric treatment on mortality in these patients. A retrospective cohort study was conducted across 13 intensive care units in Canada. We

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defined inadequate initial empiric treatment as not receiving at least one dose of an antimicrobial to which the causative pathogen(s) was susceptible within one day of initial blood culture. We evaluated the association between inadequate initial treatment and hospital mortality using a random effects multivariable logistic regression model. Among 1,190 patients (1,097 had bacteremia and 93 had candidemia), 476 (40%) died and 266 (22%) received inadequate initial treatment. Candidemic patients more often had inadequate initial empiric therapy (64.5% versus 18.8%), as well as longer delays to final culture results (4 vs 3 days) and appropriate therapy (2 vs 0 days). After adjustment, there was no detectable association between inadequate initial treatment and mortality among bacteremic patients (Odds Ratio (OR): 1.02, 95% Confidence Interval (CI) 0.70–1.48); however, candidemic patients receiving inadequate treatment had nearly three times the odds of death (OR: 2.89, 95% CI: 1.05–7.99). Inadequate initial empiric antimicrobial treatment was not associated with increased mortality in bacteremic patients, but was an important risk factor in the subgroup of candidemic patients. Further research is warranted to improve early diagnostic and risk prediction methods in candidemic patients.

Introduction

Bloodstream infections (BSI) are associated with considerable morbidity and mortality, with an estimated burden of 575,000–677,000 total episodes and 79,000–94,000 deaths per year in North America [1]. Although BSIs can be community-acquired, many infections originate in intensive care units (ICUs) because of the frequent occurrence of device-associated infections such as ventilator-associated pneumonia, central line-associated BSI and urinary tract infections [2]. A global point prevalence study performed in 75 countries found that 15% of patients in 1,265 ICUs had documented BSIs [3]. Patients who have ICU-acquired BSIs have a 3-fold higher mortality than ICU patients who do not have BSIs [4]; the attributable cost of these infections is approximately \$25,155 CAD per patient in survivors [5].

Although some studies have shown that adequate initial empiric antimicrobial therapy improves the prognosis of critically ill patients who have BSIs, others have detected no such association (reviewed in Ramphal [6]). These conflicting findings have prompted a systematic review on the methods used to assess this relationship with the goal of providing recommendations to improve the internal and external validity of future studies [7].

One strategy to achieve early adequate empiric therapy is the initial use of broad-spectrum antibiotics; however, this strategy may worsen already high rates of antimicrobial resistance [8]. Antimicrobial resistance is even more pronounced in the ICU where broad-spectrum antimicrobial agents are more commonly used for empiric treatment due to increased illness severity and risk of transmitting resistant bacteria among patients [9]. Eventually, antimicrobial resistance limits treatment options and further delays average times to effective therapy [7,10].

In light of conflicting evidence, we sought to determine the effect of inadequate initial empiric antimicrobial treatment on in-hospital mortality in a Canadian cohort of critically ill patients with BSIs. We hypothesized that inadequate initial empiric antimicrobial treatment would be associated with an increased likelihood of hospital mortality. A better understanding of the relationship between inadequate initial empiric treatment and clinical outcomes can help to inform optimal strategies to improve the prognosis of patients with severe BSIs.

Materials and Methods

Design, Setting and Population

This was a secondary analysis of the BALANCE multi-site retrospective cohort study, conducted in 13 intensive care units (ICUs) across Canada [11]. Ethical approval was provided by the Research Ethics Boards of all participating hospitals (Sunnybrook Health Sciences Center Research Ethics Board, Toronto, Ontario (ON); Research Ethics Office, University of Alberta, Edmonton, Alberta (AB); Providence Health Care Research Ethics Board, University of British Columbia (BC), Vancouver, BC; Capital Health Research Ethics Board, Halifax, Nova Scotia; University of Manitoba Health Research Ethics Board, Winnipeg, Manitoba; Comité d'éthique de la recherche (CÉR) du CHUS, Sherbrooke, Québec (QC); University of Laval Research Ethics Board, Québec, QC; Research Ethics Office, St. Michael's Hospital, Toronto, ON; Research Ethics Board, University of Western Ontario, London, ON; Ottawa Hospital Research Ethics Boards, Ottawa, ON; Research Ethics Board, Queen's University, Kingston, ON; and Conjoint Health Research Ethics Board, University of Calgary, Calgary, AB). Informed consent was waived by the research ethics boards of all participating hospitals given the retrospective study design. The cohort was accrued by looking back from December 2013 to identify the most recent consecutive critically ill patients with bloodstream infection (up to a maximum of 100 patients per ICU). Patients were eligible for the study provided that they had a blood culture positive for a pathogenic organism during their ICU admission. Patients were excluded if they had been previously enrolled, had single positive cultures with common contaminants (coagulase negative staphylococci, *Corynebacterium spp.*, *Bacillus spp.*, *Propionibacterium spp.*, *Aerococcus spp.*, *Micrococcus spp.*), or a deep-seated infection requiring extended treatment (endocarditis, osteomyelitis, septic arthritis, undrained abscess or unremoved prosthetic material) [12–15].

Data Collection and Entry

Patient demographics, reasons for admission, severity of illness, comorbidities, source of bacteremia, pathogen(s) and susceptibility, antimicrobial treatment, and clinical outcomes were abstracted by previously trained Canadian Critical Care Trials Group-affiliated research coordinators at each ICU. Data were entered into a web-based, secure electronic case report form. All patient data were anonymized and de-identified prior to analysis.

Measures

Exposure. A patient was classified as having inadequate initial empiric antimicrobial treatment if they did not receive at least one dose of an antimicrobial to which their causative pathogen(s) was susceptible within one calendar day of culture collection. The time of blood culture collection was selected as the most objective and consistent measure of suspected onset of sepsis; onset of hypotension or organ failure could not be used as an anchor because many of the patients were already critically ill in ICU prior to the acquisition of their BSI. Although dose, interval, route and therapeutic drug levels are important components of adequate antimicrobial treatment, these concepts are more applicable to defining adequacy of an overall course of antimicrobial treatment, and cannot be easily incorporated into a definition of timing of initiation of adequate treatment. Treatment adequacy was adjudicated by an infectious diseases physician blinded to the patient's clinical outcome. For patients with polymicrobial BSIs, all pathogens were required to be susceptible to the antimicrobial(s) in the regimen, but in mixed cultures the susceptibility profile of probable contaminant species (such as *Bacillus spp.*) was ignored. For the first 100 patients, a second infectious diseases physician conducted duplicate

independent adjudication, blinded to the first adjudicator's assessment. Excellent agreement on the exact start date of adequate treatment (percent agreement = 94%) was found.

Outcome. The primary outcome was all-cause in-hospital mortality.

Covariates. Covariates considered in the analysis were age, sex, body mass index, causative pathogen(s) (grouped into 11 genus categories, see [Table 1](#)), admission category (medical, surgical, trauma, burns and neurological), ICU admission due to septic shock, vasopressor use at index blood culture, severity of illness as measured by the patient's baseline Acute Physiology and Chronic Health Evaluation (APACHE) II score (measured within 24 hours of ICU admission) [16], and 12 comorbid conditions listed in [Table 1](#). The setting in which the infection was acquired was assigned as *community* if it was diagnosed on a blood culture obtained within 48 hours of hospital admission, *hospital* if the culture was obtained more than 48 hours after hospital admission, and *ICU* if it was obtained more than 48 hours after ICU admission. The patient's source of infection (vascular catheter, pneumonia/respiratory, urinary, intra-abdominal, hepato-biliary, skin and/or soft tissue, other, or unknown) was also included, based on a review of history, physical, laboratory findings, and clinician notes. Double adjudication of the first 100 patient charts indicated moderate to high agreement on source of bacteremia. Lastly, a patient was classified as having a highly resistant organism if infected with methicillin-resistant *Staphylococcus aureus*, vancomycin-resistant *Enterococci* spp, penicillin-resistant *Streptococcus pneumoniae*, extended spectrum beta-lactamase (ESBL) producing Enterobacteriaceae, carbapenem-resistant Enterobacteriaceae, carbapenem-resistant *Acinetobacter* spp; Enterobacteriaceae resistant to at least two of fluoroquinolones, aminoglycosides or trimethoprim-sulfamethoxazole; or *Acinetobacter* spp resistant to at least two of fluoroquinolones, aminoglycosides or ceftazidime [17].

Statistical Analysis

We used a random effects logistic regression model that explicitly models between- and within-ICU variation in patient populations and treatment practices to account for the hierarchical data structure. Associations between potential confounders and the exposure (treatment inadequacy) and outcome (death) were explored through univariable analyses using Pearson's χ^2 test for categorical covariates and Student's t-test for continuous covariates. Variables associated with the exposure and outcome at the $P \leq 0.20$ level were assessed for inclusion in the multivariable model using a forward fitting approach. Variables were added in order of largest to smallest effect size with the outcome; variables were retained if they modified the beta coefficient for the effect of treatment inadequacy on mortality by $\geq 10\%$ [18]. Model goodness-of-fit was assessed using the Akaike information criterion (AIC) and the Bayesian information criterion (BIC). Following variable selection, effect modification by genus of the causative organism (s) was tested through the pre-specified addition of interaction terms to the model; statistical significance was assessed using a likelihood ratio test.

Sensitivity analyses were done to examine the robustness of study results by: (1) excluding patients who died within 3 days of initial blood culture, to assess for survivor bias as patients had to survive long enough to get adequate treatment; (2) using a two-day, rather than a one-day, window for defining treatment inadequacy; and, (3) excluding patients who received no antimicrobial treatment at all, as these patients may have been assessed as not requiring treatment, or may have died too early to receive (adequate or inadequate) empiric treatment. Similarly, in a post-hoc sensitivity analysis, patients receiving less than one day of adequate treatment were excluded.

Finally, to examine whether there was a mortality gradient with increasing time to adequate antimicrobial treatment, we repeated the model building process with time to adequate

Table 1. Baseline characteristics of critically ill patients with bloodstream infections, overall and by whether the patient received inadequate initial antimicrobial treatment.

| Characteristics | All | Inadequate Initial Treatment | Adequate Initial Treatment | P-value |
|---|-------------|------------------------------|----------------------------|---------|
| | (n = 1,190) | (n = 266) | (n = 924) | |
| Age, yr (mean ± SD) | 60.2 ± 16.9 | 60.3 ± 16.5 | 60.2 ± 17.0 | 0.940 |
| Male sex, n (%) | 739 (62.1) | 158 (59.4) | 581 (62.9) | 0.303 |
| BMI, kg/m ² (mean ± SD) ^a | 28.2 ± 8.0 | 28.9 ± 8.9 | 28.0 ± 7.7 | 0.163 |
| APACHE II score (mean ± SD) ^b | 22.7 ± 8.7 | 23.5 ± 8.5 | 22.5 ± 8.7 | 0.082 |
| Admission category, n (%) | | | | |
| Medical | 925 (77.7) | 191 (71.8) | 734 (79.4) | 0.027 |
| Surgical | 133 (11.2) | 42 (15.8) | 91 (9.8) | |
| Trauma | 71 (6.0) | 20 (7.5) | 51 (5.5) | |
| Burns | 25 (2.1) | 8 (3.0) | 17 (1.8) | |
| Neurological | 33 (2.8) | 5 (1.9) | 28 (3.0) | |
| Other | 3 (0.3) | 0 (0.0) | 3 (0.3) | |
| Comorbid condition, n (%) | | | | |
| Diabetes (type 1 and 2) | 302 (25.4) | 70 (26.3) | 232 (25.1) | 0.690 |
| Congestive heart failure | 134 (11.3) | 35 (13.2) | 99 (10.7) | 0.267 |
| Chronic renal failure ^c | 74 (6.3) | 19 (7.3) | 55 (6.1) | 0.456 |
| Cirrhosis ^c | 94 (8.1) | 32 (12.4) | 62 (6.8) | 0.004 |
| Hematological malignancy ^c | 74 (6.3) | 12 (4.6) | 62 (6.8) | 0.201 |
| Solid organ malignancy | 207 (17.4) | 41 (15.4) | 166 (18.0) | 0.333 |
| Immunosuppressive therapy | 183 (15.4) | 34 (12.8) | 149 (16.1) | 0.183 |
| Chemotherapy ^c | 70 (6.0) | 14 (5.4) | 56 (6.2) | 0.649 |
| Other immunosuppressant ^c | 67 (5.7) | 9 (3.5) | 58 (6.4) | 0.075 |
| Cerebrovascular disease ^c | 76 (6.5) | 21 (8.1) | 55 (6.1) | 0.238 |
| Peripheral vascular disease ^c | 108 (9.3) | 30 (11.6) | 78 (8.6) | 0.143 |
| Obesity | 495 (41.6) | 117 (44.0) | 378 (40.9) | 0.370 |
| Acquisition of infection, n (%) | | | | |
| Community acquired | 601 (50.5) | 80 (30.1) | 521 (56.4) | <0.001 |
| Hospital acquired | 213 (17.9) | 41 (15.4) | 172 (18.6) | |
| ICU acquired | 376 (31.6) | 145 (54.5) | 231 (25.0) | |
| Highly resistant organism(s), n(%) ^d | 143 (12.0) | 53 (19.9) | 90 (9.7) | <0.001 |
| Polymicrobial infection, n(%) | 176 (14.8) | 40 (15.0) | 136 (14.7) | 0.897 |
| Genus Group | | | | |
| <i>Escherichia coli</i> | 216 (18.2) | 23 (8.6) | 193 (20.9) | <0.001 |
| <i>Staphylococcus aureus</i> | 174 (14.6) | 32 (12.0) | 142 (15.4) | 0.175 |
| <i>Enterococcus</i> spp | 148 (12.4) | 52 (19.5) | 96 (10.4) | <0.001 |
| <i>Coagulase negative staphylococci</i> | 114 (9.6) | 37 (13.9) | 77 (8.3) | 0.006 |
| <i>Klebsiella</i> spp | 108 (9.1) | 7 (2.6) | 101 (10.9) | <0.001 |
| <i>Candida</i> spp | 93 (7.8) | 60 (22.6) | 33 (3.6) | <0.001 |
| <i>Streptococcus pneumoniae</i> | 85 (7.1) | 1 (0.4) | 84 (9.1) | <0.001 |
| <i>Pseudomonas aeruginosa</i> | 69 (5.8) | 16 (6.0) | 53 (5.7) | 0.864 |
| <i>Enterobacter</i> spp | 51 (4.3) | 16 (6.0) | 35 (3.8) | 0.114 |
| <i>Alpha hemolytic streptococci</i> | 46 (3.9) | 6 (2.3) | 40 (4.3) | 0.122 |
| Other | 266 (22.4) | 57 (21.4) | 209 (22.6) | 0.681 |
| Source of infection, n (%) | | | | |
| Pneumonia | 453 (38.1) | 91 (34.2) | 362 (39.2) | 0.142 |
| Urinary tract | 241 (20.3) | 32 (12.0) | 209 (22.6) | <0.001 |
| Vascular catheter ^e | 234 (19.8) | 66 (24.9) | 168 (18.3) | 0.017 |
| Intra-abdominal | 188 (15.8) | 36 (13.5) | 152 (16.5) | 0.250 |

(Continued)

Table 1. (Continued)

| Characteristics | All | Inadequate Initial Treatment | Adequate Initial Treatment | P-value |
|--|-------------|------------------------------|----------------------------|---------|
| | (n = 1,190) | (n = 266) | (n = 924) | |
| Skin & soft tissue | 96 (8.1) | 20 (7.5) | 76 (8.2) | 0.709 |
| Hepato-biliary | 77 (6.5) | 12 (4.5) | 65 (7.0) | 0.140 |
| Other | 62 (5.2) | 11 (4.1) | 51 (5.5) | 0.371 |
| Unknown | 183 (15.4) | 71 (26.7) | 112 (12.1) | <0.001 |
| Admitted with septic shock, n(%) | 442 (37.1) | 72 (27.1) | 370 (40.0) | <0.001 |
| Vasopressor use (day 0), n(%) ^f | 602 (50.7) | 119 (44.7) | 483 (52.4) | 0.027 |
| Death, n(%) | 476 (40.0) | 135 (50.8) | 341 (36.9) | <0.001 |
| Participating ICU | | | | |
| ICU 1 | 99 (8.3) | 25 (9.4) | 74 (8.0) | <0.001 |
| ICU 2 | 82 (6.9) | 17 (6.4) | 65 (7.0) | |
| ICU 3 | 100 (8.4) | 7 (2.6) | 93 (10.1) | |
| ICU 4 | 100 (8.4) | 46 (17.3) | 54 (5.8) | |
| ICU 5 | 78 (6.6) | 16 (6.0) | 62 (6.7) | |
| ICU 6 | 100 (8.4) | 21 (7.9) | 79 (8.5) | |
| ICU 7 | 94 (7.9) | 41 (15.4) | 53 (5.7) | |
| ICU 8 | 99 (8.3) | 22 (8.3) | 77 (8.3) | |
| ICU 9 | 100 (8.4) | 6 (2.3) | 94 (10.2) | |
| ICU 10 | 100 (8.4) | 12 (4.5) | 88 (9.5) | |
| ICU 11 | 100 (8.4) | 20 (7.5) | 80 (8.7) | |
| ICU 12 | 38 (3.2) | 9 (3.4) | 29 (3.1) | |
| ICU 13 | 100 (8.4) | 24 (9.0) | 76 (8.2) | |

BMI—body mass index; APACHE—Acute Physiology and Chronic Health Evaluation; ICU—intensive care unit

^aExcludes 151 patients with missing BMI

^bExcludes 15 patients with missing APACHE II score

^cExcludes 23 patients with missing data

^dIncludes methicillin-resistant *Staphylococcus aureus*, vancomycin-resistant *Enterococci* spp, penicillin-resistant *Streptococcus pneumoniae*, extended spectrum beta-lactamase (ESBL) producing *Enterobacteriaceae*, carbapenem-resistant *Enterobacteriaceae*, carbapenem-resistant *Acinetobacter* spp; or *Enterobacteriaceae* resistant to at least two of fluoroquinolones, aminoglycosides or trimethoprim-sulfamethoxazole; or *Acinetobacter* spp resistant to at least two of fluoroquinolones, aminoglycosides or ceftazidime

^eExcludes 6 patients with missing data

^fExcludes 3 patients with missing data

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treatment in days as an ordinal exposure variable, divided into 4 categories: 0 days (no delay), 1 day (minor delay), 2–3 days (moderate delay) and ≥4 days (major delay). Patients who did not receive any adequate antimicrobial treatment were classified as experiencing a major delay in treatment.

Analyses were conducted using Stata v12 (StataCorp. College Station, TX).

Results

Patient Characteristics

The cohort included 1,190 critically ill patients with bloodstream infections. Over three quarters (77.7%) of patients were admitted to the ICU with a medical cause; one third were

specifically admitted due to septic shock (37.1%) (Table 1). Overall, 176 patients (14.8%) patients had polymicrobial BSIs. *E. coli* and *S. aureus* were the most common causative pathogens, being cultured in one third of patients. Half of the infections were acquired in the community (50.5%), with the remainder acquired in hospital or ICU. Pneumonia (38.1%), urinary tract (20.3%), and vascular catheter infections (19.8%) were the three most common sources of infection; the source of infection was unknown in 15.4% of patients. A highly resistant pathogen was cultured for 12.0% of patients. A total of 476 (40.0%) patients died in hospital.

Among all patients with bloodstream infections, 266 (22.4%, 95% CI 20.0%– 24.8%) received inadequate initial empiric antimicrobial treatment. There was wide variation across ICUs. Two ICUs (#4 and 7) that comprised one-sixth of the study’s population (16.3%) accounted for one-third of patients with inadequate initial empiric treatment (32.7%) (Table 1). In these ICUs, the proportion of patients with inadequate treatment reached as high as 46.0% compared to as low as 6.0% in other ICUs (#3 and 9) ($P < 0.001$, data not shown). Patients receiving inadequate initial empiric treatment were more likely to be admitted after surgery or have a vascular catheter identified as the source for infection (Table 1). Patients with inadequate initial empiric antimicrobial treatment were also more likely to have acquired their infection in the ICU and to have been infected with *Enterococcus* spp or *Candida* spp., as compared to those with adequate treatment.

Multivariable Analysis

In the final multivariable model, we did not detect an independent association between inadequate treatment and mortality in patients with bacteremia (adjusted OR = 1.02, 95% CI 0.70–1.48). However, in patients with candidemia, inadequate initial empiric treatment was associated with almost three times the odds of death (adjusted OR = 2.89, 95% CI 1.05–7.99) (Table 2). Interaction terms for all other genus variables were non-significant when added to the fully adjusted model (S1 Table).

Effect of time to adequate treatment on mortality

When we examined time to adequate treatment in days as the exposure variable, we found similar results. In the adjusted model, each one category increase in delay (no delay, minor, moderate and major) was not associated with increased mortality in patients with bacteremia (adjusted OR = 1.03, 95% CI = 0.89–1.21); however, a 67% increase in mortality was detected with each additional stage of delay in initiation of adequate antifungal treatment in patients with candidemia (adjusted OR = 1.67, 95% CI 1.06–2.63).

Table 2. Multivariable model results, stratified by bacteremia versus candidemia, for the effect of inadequate initial empiric treatment on patient mortality.

| Model | Stratum | OR (95% CI) | P value |
|-----------------------|------------|------------------|---------|
| Unadjusted | Bacteremia | 1.16 (0.83–1.61) | 0.379 |
| N = 1,190 | Candidemia | 3.69 (1.46–9.31) | 0.006 |
| Adjusted ^a | Bacteremia | 1.02 (0.70–1.48) | 0.934 |
| N = 1,161 | Candidemia | 2.89 (1.05–7.99) | 0.040 |

OR—odds ratio. CI—confidence interval.

^aAdjusted for admission category, vasopressor use, acquisition (community, hospital, ICU), unknown infection source, peripheral vascular disease, cirrhosis, highly resistant organism, sex and age. Age was included in the model as a continuous variable.

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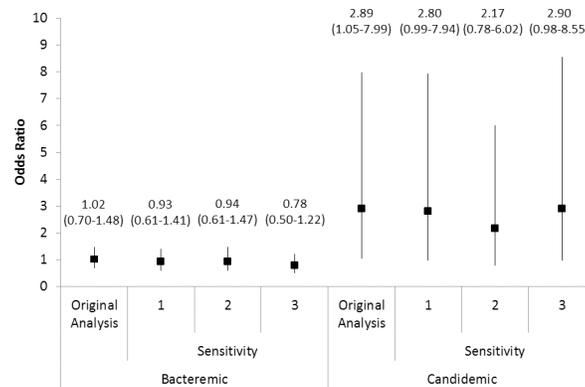


Fig 1. Adjusted odds ratios with 95% confidence intervals showing the effect of inadequate treatment on patient mortality, stratified by bacteremia versus candidemia, and under three different sensitivity analyses [1: excluding patients with early deaths (n = 1,042), 2: defining treatment inadequacy using a 2-day rather than a 1-day window (n = 1,161), 3: excluding patients who received no antimicrobial treatment (n = 1,081)].

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Sensitivity Analyses

Sensitivity analyses confirmed that our study findings were robust to excluding patients who died <3 days from initial blood collection, re-defining the initial empiric treatment window to two days, and excluding patients who received no antimicrobial treatment at all (Fig 1). Our findings were also robust to excluding 40 patients who were defined as receiving adequate treatment but for a duration of less than 1 day (bacteremic patients OR = 1.10, 95% CI = 0.75–1.61; candidemic patients OR = 2.74, 95% CI 0.98–7.68).

When we excluded patients infected with coagulase negative staphylococci there was still no association between inadequate treatment on mortality in patients with bacteremia (OR = 0.91, 95% CI 0.60–1.39). When we forced the APACHE II score into the model, results were unchanged from our main analysis, in that there was no significant association of initial inadequate treatment with bacteremia mortality (OR = 1.01, 95%CI 0.69–1.48), and a persistent association of inadequate treatment with candidemia mortality (OR = 3.08, 95%CI 1.10–8.60).

Comparing bacteremic and candidemic patients

To examine contributing factors to the observed effect modification by *Candida* spp infection, we explored differences between bacteremic and candidemic patients. Candidemic patients were more likely to be admitted to the ICU after surgery, have acquired their infection in the ICU, and have a vascular catheter or intra-abdominal source of infection (Table 3). A post-hoc analysis in which we forced solid organ malignancy, as well as vascular catheter and intra-abdominal sources of infection into the model to adjust for noted differences between bacteremic and candidemic patients (that were not already included in the main model), the findings were unchanged from our main analysis (bacteremic patients: OR 1.02, 95% CI 0.70–1.50; candidemic patients: OR 3.02, 95% CI 1.07–8.49).

Finally, there were notable differences in time to adequate treatment and receipt of final blood culture results (Fig 2). Overall, patients had a median time to treatment of 0 days (inter-quartile range (IQR) 0–0 days). Bacteremic patients received adequate treatment more promptly than candidemic patients (median 0 days, IQR 0–1 day versus median 2 days, IQR 0–3 days, $P < 0.001$). Similarly, the period from blood culture collection to receipt of final blood culture results was shorter for patients with bacteremia (median = 3.0 days, IQR 2–4) than candidemia (median = 4.0 days, IQR 2–6) ($P < 0.001$) (Fig 3).

Discussion

In our cohort of critically ill, bacteremic patients, 1 in 5 patients experienced a delay in initial adequate antimicrobial treatment, but we did not detect an overall association of inadequate treatment with hospital mortality, similar to other studies which have adjusted for key confounding variables [19–21]. Such adjustment is important in studies of heterogeneous patient populations who have variable mortality risks [6]. Our results were robust to sensitivity analyses excluding patients with early deaths, varying the definition of initial inadequate treatment, and excluding those who received no antimicrobial treatment.

However, in the subgroup of patients with candidemia, we found almost a 3-fold increased odds of hospital mortality among patients receiving inadequate initial empiric antimicrobial treatment. This finding is supported by a number of studies showing an increased risk of mortality in candidemic patients lacking adequate initial treatment [22–26]. Fungal BSIs have among the highest rates of inadequate empiric treatment, attributable mostly to a lack of empiric antifungal therapy rather than to antifungal resistance [27,28]. In our study, most of the candidemic patients (65%) did not receive adequate initial empiric treatment, and many (30%) received no adequate antimicrobial treatment. Candidemic patients experienced important delays in timing to blood culture finalization and adequate treatment provision relative to bacteremic patients. Blood culture methods have long turn-around-times to positive results, species identification and susceptibility results for *Candida* that prohibit timely diagnosis and delays initiation of antifungal therapy [29]. For each additional day that empirical fluconazole treatment is delayed, Garey *et al* (2006) detected a corresponding increase in mortality in hospitalized candidemic patients [30]. These findings have been replicated in other settings [31,32], and our own data which indicated a 67% increase in mortality with each additional stage of delayed initiation of adequate antifungal treatment.

We hypothesize that this delay in initiation or correction of therapy may have led to the observed association between inadequate initial empiric antimicrobial treatment and mortality in candidemic patients. By contrast, the shorter turn-around-time to establish diagnosis in bacteremic patients may have allowed therapy to be corrected more quickly, preventing progression to more severe illness. This notion is supported by a quasi-experimental study of surgical critically ill patients that demonstrated that taking a more conservative approach to antimicrobial treatment (i.e. delaying treatment until microbiological evidence of infection) did not worsen mortality [33]. Alternatively, it is possible that we failed to detect an association between inadequate treatment and mortality in bacteremic patients because the crucial period of reversibility may be in the first few hours from onset of infection, and we were limited to measuring timing of onset from culture collection, and also to measuring delay in calendar days rather than hours [31].

A prior randomized controlled trial in 270 ICU patients with persistent fever despite broad spectrum antibiotics detected no benefit of empiric fluconazole versus placebo [34], and so broad use of empiric fluconazole may not be warranted in all ICU patients with suspected bloodstream infection when only a minority will have candidemia (e.g. 7.8% in our cohort). However, our findings emphasize the need for strategies that minimize delay in appropriate treatment for the subset of critically ill patients at highest risk of candidemia.

Polymerase chain reaction (PCR) methods have been shown to have good sensitivity (95%) and specificity (92%) in patients with suspected invasive candidiasis relative to patients with proven candidemia [35], as well as a shorter time to initiation of antifungal treatment compared to conventional culture methods (median time: 31.0 hours vs 67.5 hours) [36]. Blood testing for 1–3 beta D glucan, a fungal cell wall component, has also been shown to expedite diagnosis of candidemia [37,38]. However, the net clinical or economic benefit of these novel

Table 3. Baseline characteristics of critically ill patients stratified by whether the patient was infected with *Candida spp.*

| Characteristics | Candida Infection | | P value |
|---|-------------------|----------------|---------|
| | Yes (n = 93) | No (n = 1,097) | |
| Age, yr (mean ± SD) | 61.3 ± 15.2 | 60.1 ± 17.1 | 0.509 |
| Male sex, n (%) | 57 (61.3) | 682 (62.2) | 0.867 |
| BMI, kg/m ² (mean ± SD) ^a | 28.0 ± 6.6 | 28.2 ± 8.1 | 0.818 |
| APACHE II score (mean ± SD) ^b | 24.3 ± 8.6 | 22.6 ± 8.7 | 0.066 |
| Admission category, n (%) | | | |
| Medical | 73 (78.5) | 852 (77.7) | 0.060 |
| Surgical | 17 (18.3) | 116 (10.6) | |
| Trauma | 2 (2.2) | 69 (6.3) | |
| Burns | 1 (1.1) | 24 (2.2) | |
| Neurological | 0 (0.0) | 33 (3.0) | |
| Other | 0 (0.0) | 3 (0.3) | |
| Comorbid condition, n (%) | | | |
| Diabetes (type 1 and 2) | 23 (24.7) | 279 (25.4) | 0.881 |
| Congestive heart failure | 14 (15.1) | 120 (10.9) | 0.228 |
| Chronic renal failure ^c | 7 (7.8) | 67 (6.2) | 0.560 |
| Cirrhosis ^c | 7 (7.8) | 87 (8.1) | 0.920 |
| Hematological malignancy ^c | 4 (4.4) | 70 (6.5) | 0.651 |
| Solid organ malignancy | 24 (25.8) | 183 (16.7) | 0.026 |
| Immunosuppressive therapy | 14 (15.1) | 169 (15.4) | 0.928 |
| Chemotherapy ^c | 6 (6.7) | 64 (5.9) | 0.781 |
| Other immunosuppressant ^c | 4 (4.4) | 63 (5.8) | 0.813 |
| Cerebrovascular disease ^c | 7 (7.8) | 69 (6.4) | 0.613 |
| Peripheral vascular disease ^c | 5 (5.6) | 103 (9.6) | 0.257 |
| Obesity | 33 (35.5) | 462 (42.1) | 0.213 |
| Acquisition of infection, n (%) | | | |
| Community acquired | 21 (22.6) | 580 (52.9) | <0.001 |
| Hospital acquired | 22 (23.7) | 191 (17.4) | |
| ICU acquired | 50 (53.8) | 326 (29.7) | |
| Highly resistant organism(s), n(%) ^d | 2 (2.2) | 141 (12.9) | 0.002 |
| Source of infection, n (%) | | | |
| Pneumonia | 42 (45.2) | 411 (37.5) | 0.142 |
| Urinary tract | 14 (15.1) | 227 (20.7) | 0.194 |
| Vascular catheter ^e | 27 (29.3) | 207 (19.0) | 0.016 |
| Intra-abdominal | 27 (29.0) | 161 (14.7) | <0.001 |
| Skin & soft tissue | 6 (6.5) | 90 (8.2) | 0.551 |
| Hepato-biliary | 4 (4.3) | 73 (6.7) | 0.511 |
| Other | 3 (3.2) | 59 (5.4) | 0.473 |
| Unknown | 14 (15.1) | 169 (15.4) | 0.928 |
| Inadequate Treatment, n(%) | 60 (64.5) | 206 (18.8) | <0.001 |
| Admitted with septic shock, n(%) | 37 (39.8) | 405 (36.9) | 0.583 |
| Vasopressor use (day 0), n(%) ^f | 57 (61.3) | 545 (49.8) | 0.034 |
| Death, n(%) | 60 (64.5) | 416 (37.9) | <0.001 |
| Participating ICU | | | |

(Continued)

Table 3. (Continued)

| Characteristics | Candida Infection | | P value |
|-----------------|-------------------|----------------|---------|
| | Yes (n = 93) | No (n = 1,097) | |
| ICU 1 | 7 (7.5) | 92 (8.4) | <0.001 |
| ICU 2 | 10 (10.8) | 72 (6.6) | |
| ICU 3 | 0 (0.0) | 100 (9.1) | |
| ICU 4 | 16 (17.2) | 84 (7.7) | |
| ICU 5 | 6 (6.5) | 72 (6.6) | |
| ICU 6 | 8 (8.6) | 92 (8.4) | |
| ICU 7 | 9 (9.7) | 85 (7.7) | |
| ICU 8 | 1 (1.1) | 98 (8.9) | |
| ICU 9 | 0 (0.0) | 100 (9.1) | |
| ICU 10 | 0 (0.0) | 100 (9.1) | |
| ICU 11 | 12 (12.9) | 88 (8.0) | |
| ICU 12 | 2 (2.2) | 36 (3.3) | |
| ICU 13 | 22 (23.7) | 78 (7.1) | |

BMI—body mass index; APACHE—Acute Physiology and Chronic Health Evaluation; ICU—intensive care unit

^aExcludes 151 patients with missing BMI

^bExcludes 15 patients with missing APACHE II score

^cExcludes 23 patients with missing data

^dIncludes methicillin-resistant *Staphylococcus aureus*, vancomycin-resistant *Enterococci* spp, penicillin-resistant *Streptococcus pneumoniae*, extended spectrum beta-lactamase (ESBL) producing *Enterobacteriaceae*, carbapenem-resistant *Enterobacteriaceae*, carbapenem-resistant *Acinetobacter* spp; or *Enterobacteriaceae* resistant to at least two of fluoroquinolones, aminoglycosides or trimethoprim-sulfamethoxazole; or *Acinetobacter* spp resistant to at least two of fluoroquinolones, aminoglycosides or ceftazidime

^eExcludes 6 patients with missing data

^fExcludes 3 patients with missing data

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methods have yet to be evaluated in a randomized controlled trial [35]. While a growing number of studies have developed tools to identify critically ill patients at high risk of candidemia [39], many have demonstrated poor validity when applied in external study populations [39]. Further work is needed to both improve the timely diagnosis of invasive *Candida* infection and to develop accurate risk prediction tools.

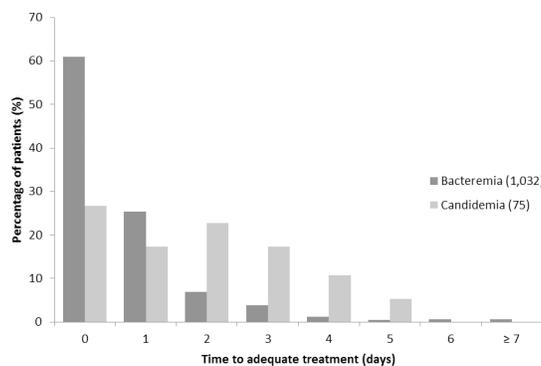


Fig 2. Time to adequate antimicrobial treatment (in days) for critically ill patients with bacteremia compared to candidemia (N = 1,107).

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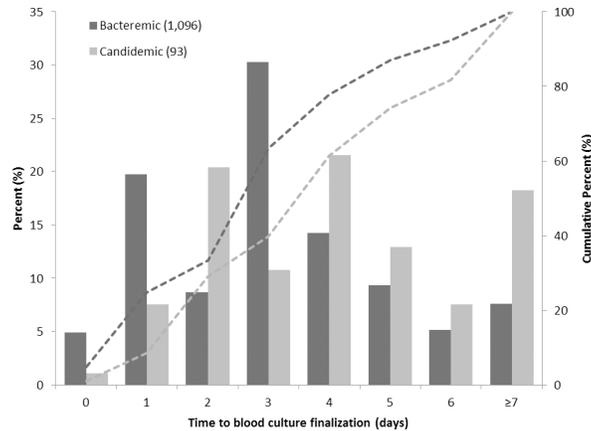


Fig 3. Time to receipt of final blood culture results (in days) for critically ill patients with bacteremia compared to candidemia.

doi:10.1371/journal.pone.0154944.g003

This study has several limitations. First, there is no standard definition of inadequate initial antimicrobial treatment [40]. As per methodologic recommendations by McGregor *et al* (2007), we defined adequacy based on whether the pathogen(s) had *in vitro* susceptibility to the administered antimicrobial(s); we did not incorporate dosing, route and clinical practice guidelines into our definition as these introduce more subjectivity [7]. Second, our estimates may be biased by our assumption that discharged patients survived; however, assuming differential misclassification where 10% of discharged patients with inadequate initial empiric antimicrobial treatment did not survive 30 days beyond discharge and 5% without inadequate initial antimicrobial treatment did not survive, the unadjusted odds ratio would have changed less than 7% (from 1.76 to 1.88). Finally, while we measured and controlled for a large number of important confounding variables (in particular, severity of illness), residual confounding by adequacy of source control and other factors remains possible.

Conclusion

In summary, while initial inadequate empiric treatment was not associated with increased mortality in our cohort of critically ill, bacteremic patients, patients with candidemia who did not receive adequate empiric therapy had a three-fold increase in the odds of death. Further work in bacteremia is needed to explain the lack of an overall association of inadequate empiric treatment and mortality by evaluating under what conditions (i.e. timing) or among which patient subgroups the effect of inadequate treatment negatively impacts patients, and also to test the safety of delaying broad-spectrum empiric antibacterial treatment in some patients. Further work in candidemia is needed to improve the timeliness of diagnosis and to develop validated risk prediction tools; both strategies have the potential to decrease delays in appropriate treatment without increasing empirical prescribing of antifungals.

Supporting Information

S1 Table. Likelihood ratio test results to evaluate whether the relationship between initial inadequate empiric antimicrobial treatment and patient mortality varied by genus of causative pathogen in patients with bloodstream infections.

(DOCX)

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Author Contributions

Conceived and designed the experiments: RAF ND. Performed the experiments: RAF AHR SMB DC PD RH AK F. Lamontagne F. Lauzier J. Marshall CMM LM J. Muscedere SR HTS ND. Analyzed the data: RDS ND. Contributed reagents/materials/analysis tools: RAF SMB DC PD RH AK F. Lamontagne F. Lauzier J. Marshall CMM LM J. Muscedere SR HTS ND. Wrote the paper: RDS RAF AHR SMB DC PD RH AK F. Lamontagne F. Lauzier J. Marshall CMM LM J. Muscedere SR HTS ND.

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